

Technical Appendix to RATIONAL AND NEAR-RATIONAL BUBBLES WITHOUT DRIFT

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Appendix: Derivations and Proofs

A.1. Separating Consumption from Dividends

The Lucas (1978) model implies $c_t = d_t$ for all t . This Section outlines a version of the model that allows $c_t \neq d_t$. The agent's first-order condition is

$$\frac{p_t}{d_t} = E_t \left[\beta \left(\frac{c_{t+1}}{c_t} \right)^{-\alpha} \left(\frac{d_{t+1}}{d_t} \right) \left(\frac{p_{t+1}}{d_{t+1}} + 1 \right) \right]. \quad (\text{A.1})$$

The separate growth rates of consumption and dividends are now given by

$$\log(c_t/c_{t-1}) \equiv x_t^c = \bar{x}^c + \rho(x_{t-1}^c - \bar{x}^c) + \varepsilon_t, \quad \varepsilon_t \sim N(0, \sigma_\varepsilon^2), \quad |\rho| < 1, \quad (\text{A.2})$$

$$\log(d_t/d_{t-1}) \equiv x_t^d = \bar{x}^d + \gamma(x_t^c - \bar{x}^c) + v_t, \quad v_t \sim N(0, \sigma_v^2), \quad (\text{A.3})$$

where v_t is uncorrelated with ε_t . As before, consumption growth is described by a univariate AR(1) process. Deviations of dividend growth from consumption growth are governed by the parameters \bar{x}^d , γ and σ_v^2 . The original Lucas model with $c_t = d_t$ can be recovered by setting $\bar{x}^d = \bar{x}^c$, $\gamma = 1$, and $\sigma_v = 0$. For the model with $c_t \neq d_t$, these parameters are calibrated to match three moments:

- (i) the unconditional mean of dividend growth $E[\log(d_t/d_{t-1})]$,
- (ii) the contemporaneous correlation between dividend growth and consumption growth $\text{corr}(x_t^d, x_t^c)$, and
- (iii) the unconditional variance of dividend growth $\text{var}(x_t^d)$.

The resulting calibration formulas are

$$\bar{x}^d = E[\log(d_t/d_{t-1})], \quad (\text{A.4})$$

$$\gamma = \text{corr}(x_t^d, x_t^c) [\text{var}(x_t^d) / \text{var}(x_t^c)]^{0.5}, \quad (\text{A.5})$$

$$\sigma_v = [\text{var}(x_t^d) - \gamma^2 \text{var}(x_t^c)]^{0.5}. \quad (\text{A.6})$$

The agent's first-order condition can be written in terms of the price–dividend ratio y_t as follows:

$$\begin{aligned} y_t &= E_t \{ \beta \exp[-\alpha x_{t+1}^c + \bar{x}^d + \gamma(x_{t+1}^c - \bar{x}^c) + v_{t+1}](y_{t+1} + 1) \}, \\ &= E_t [\tilde{\beta} \exp(\tilde{\theta} \tilde{x}_{t+1})(y_{t+1} + 1)], \end{aligned} \quad (\text{A.7})$$

where

$$\tilde{\beta} \equiv \beta \exp(\bar{x}^d - \gamma \bar{x}^c), \quad \tilde{\theta} \equiv \gamma - \alpha, \quad \tilde{x}_t \equiv x_t^c + v_t / \tilde{\theta}.$$

Making use of the above definitions, (A.2) and (A.3) yield the following transformed version of (3):

$$\tilde{x}_t = \bar{x}^c + \rho(\tilde{x}_{t-1} - \bar{x}^c) + \omega_t, \quad \omega_t \equiv \varepsilon_t + (v_t - \rho v_{t-1}) / \tilde{\theta}, \quad (\text{A.8})$$

where $\omega_t \sim N(0, \sigma_\omega^2)$, $\sigma_\omega^2 = \sigma_\varepsilon^2 + (1 + \rho^2)\sigma_v^2/\tilde{\theta}^2$, and $\text{corr}(\omega_t, \omega_{t-1}) = -\rho/(\tilde{\theta}^2\sigma_\varepsilon^2/\sigma_v^2 + 1 + \rho^2)$.

Finally, we define $\tilde{z}_t \equiv \tilde{\beta} \exp(\tilde{\theta} \tilde{x}_t)(y_t + 1)$ to obtain the following transformed version of (8):

$$\tilde{z}_t = \tilde{\beta} \exp(\tilde{\theta} \tilde{x}_t)(E_t \tilde{z}_{t+1} + 1). \quad (\text{A.9})$$

Thus, by an appropriate change of variables, (A.8) and (A.9) retain the same basic forms as (3) and (8), with the exception that the innovation ω_t is not iid but instead exhibits serial correlation. However, for small values of ρ , I can make the simplifying assumption that $\text{corr}(\omega_t, \omega_{t-1}) \simeq 0$. With this assumption, all of the article's theoretical results will go through when expressed in terms of the transformed variables.

A.2. Proof of Proposition 1: Approximate Fundamental Solution

Iterating ahead the conjectured law of motion for z_t^f and taking the conditional expectation yields

$$E_t z_{t+1}^f = \exp[a_0 + \rho a_1(x_t - \bar{x}) + \frac{1}{2}(a_1)^2 \sigma_\varepsilon^2]. \quad (\text{A.10})$$

Substituting the above expression into the first-order condition (8) and then taking logarithms yields

$$\begin{aligned} \log(z_t^f) &= F(x_t) = \log(\beta) + \theta x_t \\ &\quad + \log\{\exp[a_0 + \rho a_1(x_t - \bar{x}) + \frac{1}{2}(a_1)^2 \sigma_\varepsilon^2] + 1\}, \\ &\simeq a_0 + a_1(x_t - \bar{x}), \end{aligned} \quad (\text{A.11})$$

where $a_0 \equiv E[\log(z_t^f)]$ and a_1 are Taylor-series coefficients which are given by

$$F(\bar{x}) = a_0 = \log(\beta) + \theta \bar{x} + \log\{\exp[a_0 + \frac{1}{2}(a_1)^2 \sigma_\varepsilon^2] + 1\} \quad (\text{A.12})$$

$$F'(\bar{x}) = a_1 = \theta + \frac{\rho a_1 \exp[a_0 + \frac{1}{2}(a_1)^2 \sigma_\varepsilon^2]}{\exp[a_0 + \frac{1}{2}(a_1)^2 \sigma_\varepsilon^2] + 1}. \quad (\text{A.13})$$

Solving (A.12) for a_0 yields

$$a_0 = \log\left\{ \frac{\beta \exp(\theta \bar{x})}{1 - \beta \exp[\theta \bar{x} + \frac{1}{2}(a_1)^2 \sigma_\varepsilon^2]} \right\}, \quad (\text{A.14})$$

which can be substituted into (A.13) to yield the following non-linear equation that determines a_1 :

$$a_1 = \theta + \rho a_1 \beta \exp[\theta \bar{x} + \frac{1}{2}(a_1)^2 \sigma_\varepsilon^2]. \quad (\text{A.15})$$

Solving (A.15) for a_1 yields the non-linear equation shown in Proposition 1. There are two solutions but only one solution satisfies the condition $\beta \exp[\theta \bar{x} + \frac{1}{2}(a_1)^2 \sigma_\varepsilon^2] < 1$ such that $\exp(a_0) = \exp[E \log(z_t^f)] > 0$.

A.3. Asset Pricing Moments: Fundamental Solution

This Section briefly outlines the derivation of (16)–(19).

Equation (16) follows directly from (15) by taking the unconditional expectation of $\log(y_t^f)$. We have

$$\log(y_t^f) - E[\log(y_t^f)] = a_1 \rho_1 (x_t - \bar{x}), \quad (\text{A.16})$$

which implies $\text{var}[\log(y_t^f)] = (a_1)^2 \rho^2 \text{var}(x_t)$, as given by (17).

The fundamental equity return can be written as

$$\begin{aligned} R_{t+1}^f &= \left(\frac{y_{t+1}^f + 1}{y_t^f} \right) \exp(x_{t+1}), \\ &= \left(\frac{z_{t+1}^f}{\beta E_t z_{t+1}^f} \right) \exp(\alpha x_{t+1}), \end{aligned} \quad (\text{A.17})$$

where I have eliminated y_t^f using the equilibrium relationship $y_t^f = E_t z_{t+1}^f$ and eliminated y_{t+1}^f using the definitional relationship $y_{t+1}^f + 1 = \beta^{-1} \exp(-\theta x_{t+1}) z_{t+1}^f$. Substituting in $z_{t+1}^f = \exp[a_0 + a_1(x_t - \bar{x})]$ from Proposition 1 and $E_t z_{t+1}^f$ from (15) and then taking the unconditional expectation of $\log(R_{t+1}^f)$ yields (18). I have

$$\log(R_{t+1}^f) - E[\log(R_{t+1}^f)] = \alpha(x_{t+1} - \bar{x}) + a_1 \varepsilon_{t+1}, \quad (\text{A.18})$$

which in turns implies

$$\text{var}[\log(R_{t+1}^f)] = \alpha^2 \text{var}(x_t) + (a_1)^2 \sigma_\varepsilon^2 + 2\alpha a_1 \text{cov}(x_t, \varepsilon_t). \quad (\text{A.19})$$

Substituting for $\text{var}(x_t)$ and $\text{cov}(x_t, \varepsilon_t)$ in the above expression yields (19).

A.4 Proof of Proposition 2: Continuum of Intrinsic Rational Bubbles

First consider the case where the agent can make use of the contemporaneous realisation z_t^b when forming the rational expectation $E_t z_{t+1}^b$. Iterating ahead the conjectured law of motion for z_t^b by one period and then taking the conditional expectation yields

$$E_t z_{t+1}^b = z_t^b \exp[\lambda_0 + (\rho \lambda_1 + \lambda_2)(x_t - \bar{x}) + \frac{1}{2}(\lambda_1)^2 \sigma_\varepsilon^2]. \quad (\text{A.20})$$

Substituting the above expression into the no-arbitrage condition (21) and then taking logarithms yields

$$0 = \log(\beta) + \theta x_t + \lambda_0 + (\rho \lambda_1 + \lambda_2)(x_t - \bar{x}) + \frac{1}{2}(\lambda_1)^2 \sigma_\varepsilon^2, \quad (\text{A.21})$$

where $\log(z_t^b)$ has been cancelled from both sides. For (A.21) to hold, the constant terms and the coefficients on x_t must separately sum to zero. Equilibrium therefore requires

$$\frac{1}{2}(\lambda_1)^2 \sigma_\varepsilon^2 - \underbrace{(\rho \lambda_1 + \lambda_2) \bar{x}}_{-\theta} + \log(\beta) + \lambda_0 = 0, \quad (\text{A.22})$$

$$\theta + \rho \lambda_1 + \lambda_2 = 0, \quad (\text{A.23})$$

which represent a system of two equations in three unknown constants λ_0 , λ_1 , and λ_2 . The solutions to (A.22) and (A.23) define a continuum of intrinsic rational bubble equilibria.

Now consider the case where the agent can only make use of the lagged realisation z_{t-1}^b when forming $E_t z_{t+1}^b$. Iterating ahead the conjectured law of motion for z_t^b by one period and then substituting out z_t^b using the same law of motion yields

$$z_{t+1}^b = z_{t-1}^b \exp\{2\lambda_0 + [\lambda_1(1 + \rho) + \lambda_2](x_t - \bar{x}) + \lambda_2(x_{t-1} - \bar{x}) + \lambda_1 \varepsilon_{t+1}\}, \quad (\text{A.24})$$

where I have eliminated $(x_{t+1} - \bar{x})$ using the law of motion for consumption/dividend growth (33). Taking the conditional expectation of the above expression yields

$$E_t z_{t+1}^b = z_{t-1}^b \exp\left\{2\lambda_0 + [\lambda_1(1 + \rho) + \lambda_2](x_t - \bar{x}) + \lambda_2(x_{t-1} - \bar{x}) + \frac{1}{2}(\lambda_1)^2 \sigma_\varepsilon^2\right\}. \quad (\text{A.25})$$

Substituting the above expression into the no-arbitrage condition (21) and then taking logarithms yields

$$\begin{aligned} \log(z_t^b) &= \log(z_{t-1}^b) + \log(\beta) + \theta x_t + 2\lambda_0 + [\lambda_1(1 + \rho) + \lambda_2](x_t - \bar{x}) \\ &\quad + \lambda_2(x_{t-1} - \bar{x}) + \frac{1}{2}(\lambda_1)^2 \sigma_\varepsilon^2, \end{aligned} \quad (\text{A.26})$$

which can be compared to the following expression for the logarithm of the conjectured law of motion:

$$\log(z_t^b) = \log(z_{t-1}^b) + \lambda_0 + \lambda_1(x_t - \bar{x}) + \lambda_2(x_{t-1} - \bar{x}). \quad (\text{A.27})$$

Equation (A.26) will coincide exactly with (A.27) when the following equilibriums conditions are satisfied

$$\log(\beta) + 2\lambda_0 - [\lambda_1(1 + \rho) + 2\lambda_2]\bar{x} + \frac{1}{2}(\lambda_1)^2 \sigma_\varepsilon^2 = \lambda_0 - (\lambda_1 + \lambda_2)\bar{x}, \quad (\text{A.28})$$

$$\theta + \lambda_1(1 + \rho) + \lambda_2 = \lambda_1, \quad (\text{A.29})$$

which are isomorphic to the equilibrium conditions (A.22) and (A.23).

A.5. Asset Pricing Moments: Near-rational Solution

Starting from the approximate ALM (36), the law of motion of $\Delta \log(z_t)$ can be written as:

$$\Delta \hat{z}_t = (k - 1)[\hat{z}_{t-1} - E(\hat{z}_t)] + m(x_t - \bar{x}), \quad (\text{A.30})$$

where $\hat{z}_t \equiv \log(z_t)$ and $\Delta \hat{z}_t \equiv \log(z_t/z_{t-1})$. The above equation implies:

$$\text{cov}(\Delta \hat{z}_t, x_t) = (k - 1)\text{cov}(\hat{z}_{t-1}, x_t) + m\text{var}(x_t). \quad (\text{A.31})$$

From (36), I have $\text{cov}(\hat{z}_{t-1}, x_t) = [\rho m / (1 - \rho k)] \text{var}(x_t)$, which can be substituted into (A.31) to yield (40) in the text.

The nonlinear ALM for the price–dividend ratio, (33), can be rewritten as follows:

$$\begin{aligned} y_t &= (y_{t-1} + 1)\beta \exp\left[b(1 + \rho)(x_t - \bar{x}) + \theta x_{t-1} + \frac{1}{2}b^2 \sigma_\varepsilon^2\right], \\ &= (y_{t-1} + 1)k \exp\left[\left(\frac{m - \theta}{k}\right)(x_t - \bar{x}) + \theta(x_{t-1} - \bar{x})\right], \end{aligned} \quad (\text{A.32})$$

where I have eliminated b and b^2 using the expressions for the Taylor series coefficients k and m , as given by (37) and (38). Taking logarithms of the above expression yields

$$\begin{aligned}\hat{y}_t &= \log[\exp(\hat{y}_{t-1}) + 1] + \log(k) + \left(\frac{m-\theta}{k}\right)(x_t - \bar{x}) + \theta(x_{t-1} - \bar{x}), \\ &\simeq n_0 + n_1[\hat{y}_{t-1} - E(\hat{y}_t)] + \left(\frac{m-\theta}{k}\right)(x_t - \bar{x}) + \theta(x_{t-1} - \bar{x}),\end{aligned}\quad (\text{A.33})$$

where $\hat{y}_t \equiv \log(y_t)$ and n_0 and n_1 are Taylor series coefficients. Straightforward computations yield $n_0 = \log[k/(1-k)]$ and $n_1 = k$. The unconditional expectation of the above expression yields $E(\hat{y}_t) = n_0$, as given by (41).

Using (A.33), the unconditional variance can be computed as follows:

$$\begin{aligned}\text{var}(\hat{y}_t) &= E\left\{[\hat{y}_t - E(\hat{y}_t)]^2\right\}, \\ &= \left(\frac{1}{1-k^2}\right)\left[\left(\frac{m-\theta}{k}\right)^2 + \theta^2 + 2\left(\frac{m-\theta}{k}\right)\theta\rho\right]\text{var}(x_t), \\ &\quad + \left[\frac{2(m-\theta)\rho + 2\theta k}{1-k^2}\right]\text{cov}(\hat{y}_t, x_t),\end{aligned}\quad (\text{A.34})$$

where $\text{cov}(\hat{y}_t, x_t)$ can also be computed from (A.33).

The equity return is given by

$$\begin{aligned}R_{t+1} &= \left(\frac{z_{t+1}}{\beta \hat{E}_t z_{t+1}}\right) \exp(\alpha x_{t+1}) \\ &= \frac{z_t^k z^{1-k} \exp[m(x_{t+1} - \bar{x}) + \alpha x_{t+1}]}{\beta z_{t-1} \exp[b(1+\rho)(x_t - \bar{x}) + \frac{1}{2}b^2\sigma_\varepsilon^2]},\end{aligned}\quad (\text{A.35})$$

where I have substituted in the approximate ALM (36) and the subjective expectation (30). Taking the unconditional expectation of $\hat{R}_{t+1} \equiv \log(R_{t+1})$ yields (42). From (A.35), I have

$$\begin{aligned}\hat{R}_{t+1} - E(\hat{R}_{t+1}) &= \underbrace{k[\hat{z}_t - E(\hat{z}_t)]}_{k^2[\hat{z}_{t-1} - E(\hat{z}_t)] + km(x_t - \bar{x})} - [\hat{z}_{t-1} - E(\hat{z}_t)] \\ &\quad + (m + \alpha)(x_{t+1} - \bar{x}) - b(1 + \rho)(x_t - \bar{x}),\end{aligned}\quad (\text{A.36})$$

which can be used to compute an analytical expression for $\text{var}(\hat{R}_{t+1})$.

From (43), the law of motion for the percentage forecast error is given by

$$\begin{aligned}err_{t+1} - E(err_{t+1}) &= -(1 - k^2)[\hat{z}_{t-1} - E(\hat{z}_t)] \\ &\quad + [km + m\rho - b(1 + \rho)](x_t - \bar{x}) + m\varepsilon_{t+1},\end{aligned}\quad (\text{A.37})$$

where I have eliminated $[\hat{z}_t - E(\hat{z}_t)]$ using the approximate ALM (36). Equation (A.37) is used to compute $\text{Corr}(err_{t+1}, err_t) = \text{Cov}(err_{t+1}, err_t) / \text{var}(err_{t+1})$ and $\text{RMSPE} = [E(err_{t+1}^2)]^{0.5}$ which are plotted in Figure 3.